# UC Berkeley <br> Department of Electrical Engineering and Computer Sciences 

## EE126: Probability and Random Process

Midterm 1 Solution
Fall 2014

Problem 1. (a) The size of the sample space is the number of different ways that 52 objects can be divided in 4 groups of 13, and is given by

$$
\frac{52!}{13!.13!.13!.13!} .
$$

There are 4! different ways of distributing the 4 aces to the 4 players, and there are

$$
\frac{48!}{12!.12!.12!.12!}
$$

different ways of dividing the remaining 48 cards into 4 groups of 12. Thus, the probability is

$$
\frac{\frac{48!4!}{\frac{12!\cdot 12!+12!12!}{5!}}}{13!\cdot 13!\cdot 13!!13!} .
$$

(b) Clearly $X$ takes value in $\{4,5,6\}$. The probability that $X=4$ is the probability that you get 4 in all 3 exams that is $\operatorname{Pr}(X=4)=(1 / 3)^{3}=1 / 27$. The probability that $X=5$ is the probability that you get 4 or 5 in all 3 exams minus the probability that you get 4 in all 3 exams that is $\operatorname{Pr}(X=5)=$ $(2 / 3)^{3}-(1 / 3)^{3}=7 / 27$. Then, $\operatorname{Pr}(X=6)=1-1 / 27-7 / 27=19 / 27$.
(c) Figure 1 shows how the Huffman coding is done. Then, $A, B$ and $E$ are encoded to 2 bits, and $R$ and $S$ are encoded to 3 bits. Thus, "Bears" is encoded to 12 bits.
(d) The clever way to solve this problem is to notice that the 4 points are dropped at random, and for any four points on the circle there are 3 ways of connecting them to make two lines and the two lines intersect in one case. See Figure 2. Thus, the probability is $1 / 3$.
(e) Let $F$ be the event that the fair coin is picked and $A$ be the event that there are 3 Heads in the 4 coin flips. We are interested in calculating $\operatorname{Pr}(F \mid A)$. By Bayes rule, we have

$$
\begin{aligned}
\operatorname{Pr}(F \mid A) & =\frac{\operatorname{Pr}(A \mid F) \operatorname{Pr}(F)}{\operatorname{Pr}(A \mid F) \operatorname{Pr}(F)+\operatorname{Pr}(A \mid \bar{F}) \operatorname{Pr}(\bar{F})} \\
& =\frac{\binom{4}{3}(1 / 2)^{4}}{\binom{4}{3}(1 / 2)^{4}+\binom{4}{3}(3 / 4)^{3}(1 / 4)}
\end{aligned}
$$



Figure 1: Huffman Coding of $(A, B, E, R, S)$.


Figure 2: 4 points on the circle

Problem 2. (a) The shaded area should integrate to 1 so $A / 8+A / 8=1$ so $A=4$.
(b) By symmetry of the pdf across the line $x=0.5$, we find that $E[X]=1 / 2$. To find $E(Y)$ again by symmetry we have

$$
E(Y)=2 \int_{y=0}^{1 / 2} \int_{x=0}^{y} A y d x d y=2 A \int_{0}^{1 / 2} y^{2}=2 A / 24=1 / 3
$$

(c) $X$ and $Y$ are not independent. As an example, given that $X=1 / 4, Y$ has to be larger than $1 / 4$. They are uncorrelated since by symmetry $E[X \mid Y=y]=$ $E[X]$. Then, $\operatorname{cov}(X, Y)=0$ as shown in Homework 3 .

Problem 3. (a) Since nobody is waiting to be served, you are the head-of-the-line customer in the queue. By memoryless property of the exponential distribution, the remaining service time of the customers getting service is again exponential with mean 1. Let $X_{1}$ be the service time of the customer at cashier 1, and $X_{2}$ be the service time of the customer at cashier 2. Let $X$ be your serving time. Then, $X_{1}, X_{2}$ and $X$ are iid exponential random variables with rate 1 . At Target pharmacy, since there is a central queue, as soon as a cashier becomes free, your service starts. Thus, your waiting time is $T_{1}=\min \left(X_{1}, X_{2}\right)+X$. At Safeway pharmacy, you have to pick a line at first. Since the service times are exponential it does not matter which line to choose. Without loss of generality suppose you stand in line 1 . Then, your
waiting time is $T_{2}=X_{1}+X$. Let's find the distribution of minimum of two exponentials. Let $Y=\min \left(X_{1}, X_{2}\right)$. Then,

$$
F_{Y}(y)=1-\operatorname{Pr}(Y \geq y)=1-\operatorname{Pr}\left(X_{1} \geq y, X_{2} \geq y\right)=1-e^{-2 y}
$$

Thus, $Y$ is exponentially distributed with mean $1 / 2$. Thus, $E\left[T_{1}\right]=1 / 2+1=$ $3 / 2$. Clearly, $E\left[T_{2}\right]=1+1=2$. Thus, the queueing strategy at Target pharmacy is better.
(b) Similarly, $T_{1}=\min \left(X_{1}, X_{2}, \ldots, X_{n}\right)+X$ and $T_{2}=X_{1}+X$. Let $Y=$ $\min \left(X_{1}, X_{2}, \ldots, X_{n}\right)$. Then,

$$
F_{Y}(y)=1-\operatorname{Pr}(Y \geq y)=1-\operatorname{Pr}\left(X_{1} \geq y, X_{2} \geq y, \ldots, X_{n} \geq y\right)=1-e^{-n y}
$$

Thus, $Y$ is exponentially distributed with mean $1 / n$. Thus, $E\left[T_{1}\right]=1 / n+1=$ $(n+1) / n$. Similar to previous part, $E\left(T_{2}\right)=2$.

Problem 4. We consider the four vertices of the square in which the center of the circle lies. For each of these vertices, there is some probability $p$ that it is in the circle. Accordingly, we can write $X=X_{1}+X_{2}+X_{3}+X_{4}$, where $X_{i}$ is 1 if vertex $i$ of the square is in the circle and is 0 otherwise. Now,

$$
E(X)=4 E\left(X_{1}\right)=4 p
$$

The key observation here is that the average value of a sum of random variables is the sum of their average values, even when these random variables are not independent. It remains to calculate $p$. To do that, note that the set of possible locations of the center of the circle in a given square such that one vertex is in the circle is a quarter-circle with radius 1. Hence, $p=\pi / 4$ and we conclude that $E(X)=\pi$.

Problem 5. (a) At each spinning of the wheel, one ball is dropped in a particular bin with probability $d / n$. Thus, the number of balls in a bin is a binomial random variable $X \sim B i(m, d / n)$. So $E(X)=m d / n$ and $\operatorname{var}(X)=\frac{m d}{n}(1-$ $d / n)$.
(b) Since $m$ is large and $m d / n=0.2$ is a constant, we use Poisson approximation. Thus, $\operatorname{Pr}(X=x) \simeq \frac{e^{-m d / n}(m d / n)^{x}}{x!}$.
(c) Note that

$$
\operatorname{Pr}\left(E_{i+1} \mid E_{i}\right)=\frac{\operatorname{Pr}\left(E_{i+1} \cap E_{i}\right)}{\operatorname{Pr}\left(E_{i}\right)}
$$

We have

$$
\begin{aligned}
\operatorname{Pr}\left(E_{i+1} \cap E_{i}\right) & =1-\operatorname{Pr}\left(\bar{E}_{i+1}\right)-\operatorname{Pr}\left(\bar{E}_{i}\right)+\operatorname{Pr}\left(\bar{E}_{i} \cap \bar{E}_{i+1}\right) \\
& =1-2\left(\left(1-\frac{d}{n}\right)^{m}\right)+\left(1-\frac{d+1}{n}\right)^{m}
\end{aligned}
$$

Then,

$$
\operatorname{Pr}\left(E_{i+1} \mid E_{i}\right)=\frac{1-2\left(\left(1-\frac{d}{n}\right)^{m}\right)+\left(1-\frac{d+1}{n}\right)^{m}}{1-\left(1-\frac{d}{n}\right)^{m}}
$$

Clearly, $\operatorname{Pr}\left(E_{i+1} \mid E_{i}\right) \neq \operatorname{Pr}\left(E_{i+1}\right)$ so they are not independent.
(d) The probability of a bin being empty is $(1-d / n)^{m}$. Thus, the expected number of empty bins is $n(1-d / n)^{m}$.
(e) We upper bound $\operatorname{Pr}(E)$ using union bound as follows. Let $A_{i}$ be the event that $\operatorname{bin} i$ is empty. Then,

$$
\operatorname{Pr}(E)=\operatorname{Pr}\left(A_{1} \cup A_{2} \ldots \cup A_{n}\right) \leq n \operatorname{Pr}\left(A_{1}\right)=n(1-d / n)^{m} .
$$

Replacing the values of $d$ and $m$ we have

$$
\lim _{n \rightarrow \infty} \operatorname{Pr}(E) \leq \lim _{n \rightarrow \infty} n(1-2 / n)^{n \ln (n)}=\lim _{n \rightarrow \infty} n e^{-2 \ln (n)}=\lim _{n \rightarrow \infty} n / n^{2}=0
$$

Thus, $\lim _{n \rightarrow \infty} \operatorname{Pr}(E)=0$.

